

Logistic Regression, Exponential Family

Junxian He Feb 9, 2024 COMP 5212 Machine Learning Lecture 3







Classification

CAT

Labels are discrete





Logistic Regression

Want $h_{\theta}(x) \in [0, 1]$. Let's pick a smooth function:

- Given a training set $\{(x^{(i)}, y^{(i)}) \text{ for } i = 1, ..., n\}$ let $y^{(i)} \in \{0, 1\}$.
 - $h_{\theta}(x) = g(\theta^T x)$

Logistic Regression

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There are many options of g....

Logistic Regression

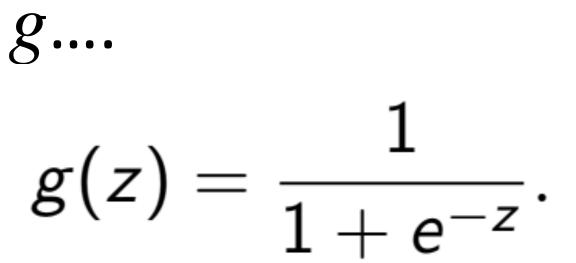
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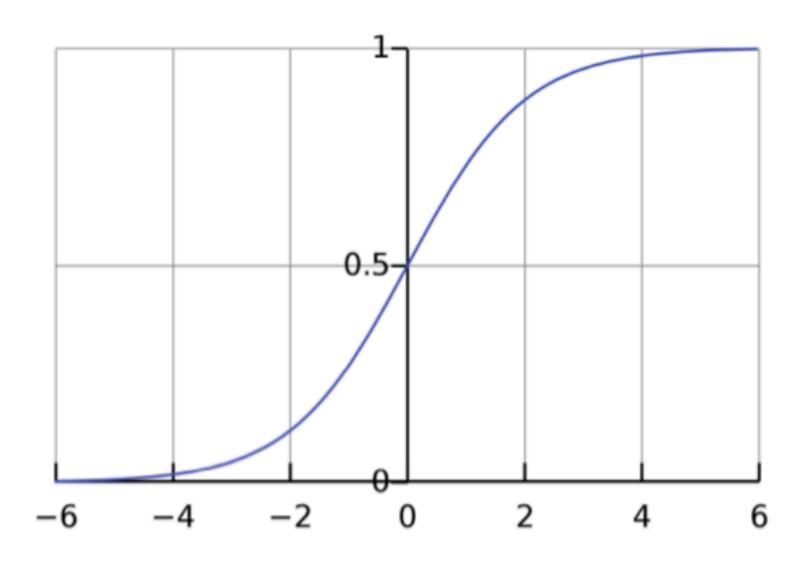
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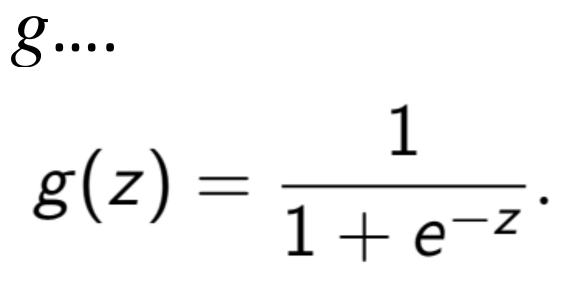
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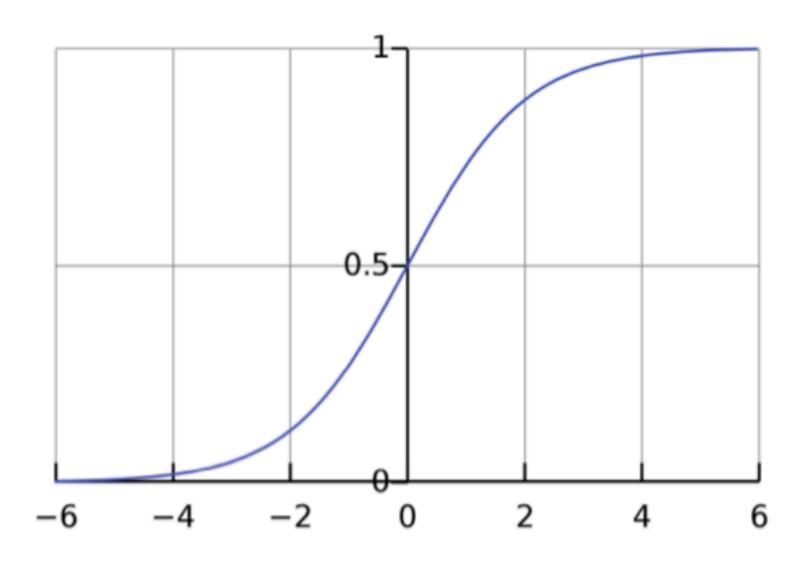
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Logistic Regression

 $h_{\theta}(x) = g(\theta^T x)$ Link Function

$$g(z) = rac{1}{1+e^{-z}}.$$

How do we interpret $h_{\theta}(x)$?

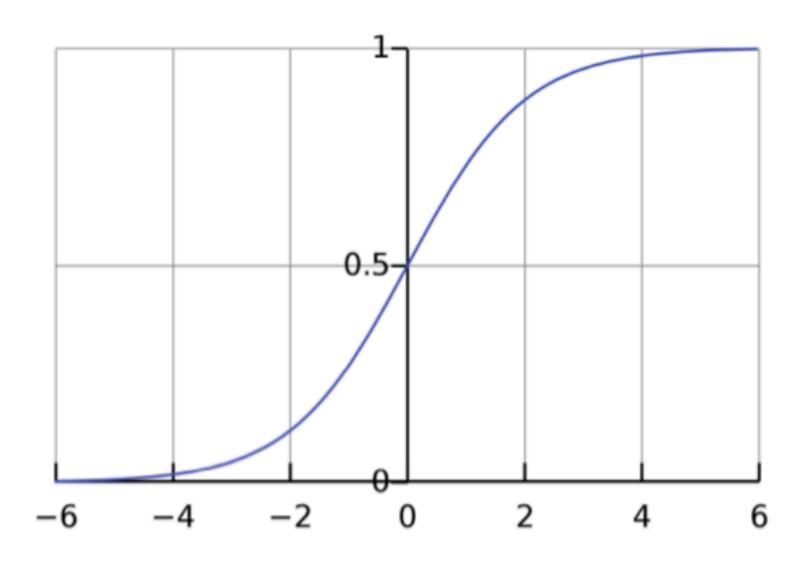
$$P(y = 1 \mid x; \theta) = h_{\theta}(x)$$

 $P(y = 0 \mid x; \theta) = 1 - h_{\theta}(x)$



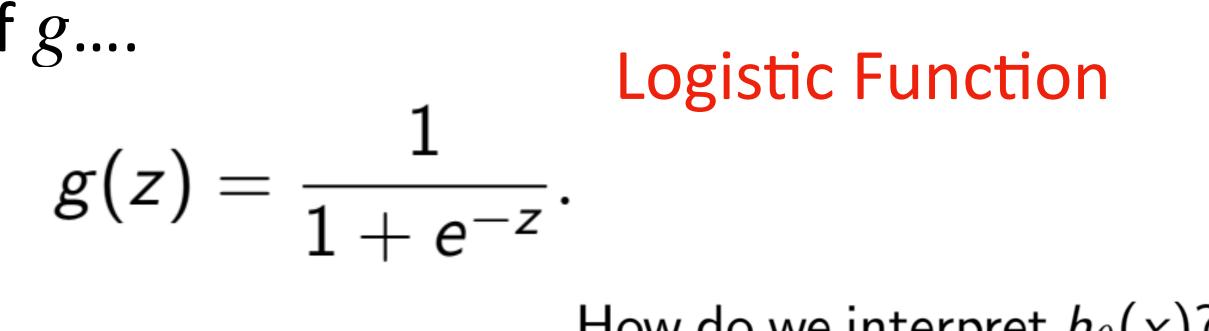
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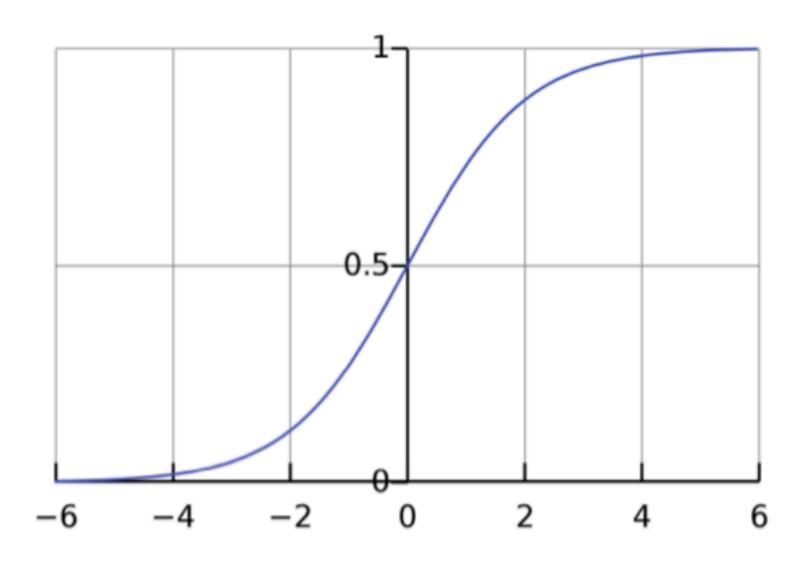
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 $h_{\theta}(x)$



$$P(y = 1 \mid x; \theta) = h_{ heta}(x \mid P(y = 0 \mid x; \theta)) = 1 - h_{ heta}(x \mid x; \theta)$$

Then,

$$L(\theta) = P(y \mid X; \theta) = \prod_{i=1}^{n} p(y^{(i)} \mid x^{(i)}; \theta)$$



 $h_{\theta}(x)$

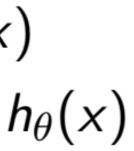


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Ne want to express "if-then" logics, how?

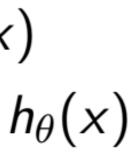


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$$L(\theta) = P(y \mid X; \theta) = \prod_{i=1}^{n} p(y^{(i)} \mid x^{(i)}; \theta) \quad \bigvee$$
$$= \prod_{i=1}^{n} h_{\theta}(x^{(i)})^{y^{(i)}}(1 - h_{\theta}(x^{(i)}))^{y^{(i)}}(1 - h_{\theta}(x^{(i)}))^{y^$$





Ve want to express "if-then" logics, how?

 $(i)))^{1-y(i)}$



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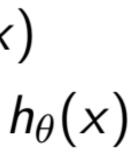
Then,

$$egin{aligned} L(heta) = & P(y \mid X; heta) = \prod_{i=1}^n p(y^{(i)} \mid x^{(i)}; heta) & V(x) \ &= & \prod_{i=1}^n h_ heta(x^{(i)})^{y^{(i)}}(1 - h_ heta(x^{(i)}))^{y^{(i)}}) & V(x) \ &= & V($$

Taking logs to compute the log likelihood $\ell(\theta)$ we have:

$$\ell(heta) = \log L(heta) = \sum_{i=1}^{n} y^{(i)} \log h_{ heta}(x^{(i)}) + (1 - y^{(i)}) \log(1 + 1)$$





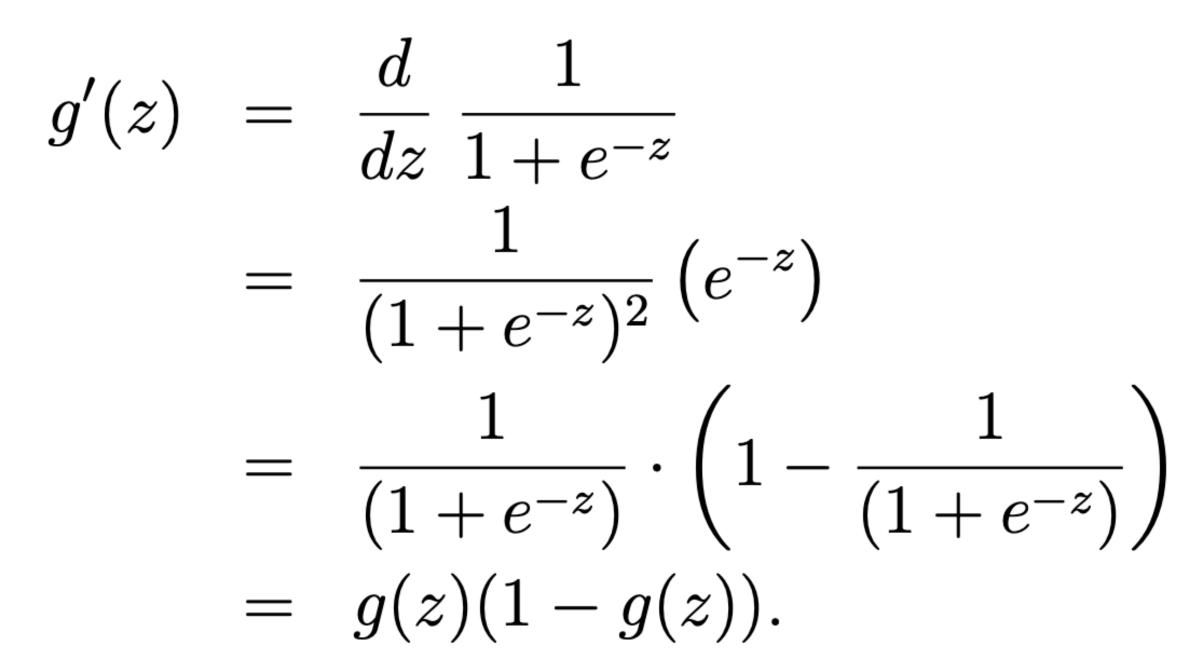
Ve want to express "if-then" logics, how?

 $(i)))^{1-y^{(i)}}$

Maximum likelihood estimation $-h_{\theta}(x^{(i)}))$



Derivative of Logistic Function



Gradient Descent

$$\begin{split} \frac{\partial}{\partial \theta_j} \ell(\theta) &= \left(y \frac{1}{g(\theta^T x)} - (1 - y) \frac{1}{1 - g(\theta^T x)} \right) \frac{\partial}{\partial \theta_j} g(\theta^T x) \\ &= \left(y \frac{1}{g(\theta^T x)} - (1 - y) \frac{1}{1 - g(\theta^T x)} \right) g(\theta^T x) (1 - g(\theta^T x)) \frac{\partial}{\partial \theta_j} \theta^T x \\ &= \left(y (1 - g(\theta^T x)) - (1 - y) g(\theta^T x) \right) x_j \\ &= \left(y - h_{\theta}(x) \right) x_j \end{split}$$

$$\theta_j := \theta_j + \alpha \left(\right.$$

 $\left(y^{(i)} - h_{\theta}(x^{(i)})\right) x_j^{(i)}$

Gradient Descent

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$$\theta_j := \theta_j + \alpha \left(y^{(i)} - h_{\theta}(x^{(i)}) \right) x_j^{(i)}$$

Looks identical to LMS update rule

IS update rule in linear regression

Gradient Descent

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$$\theta_j := \theta_j + \alpha \left(y^{(i)} - h_\theta(x^{(i)}) \right) x_j^{(i)}$$

Looks identical to LMS update rule in linear regression Is this coincidence?



{Cat, dog, dragon, fish, pig}

Given a training set $\{(x^{(1)}, y^{(1)}), \dots, (x^{(n)}, y^{(n)})\}, y^{(i)} \in \{1, 2, \dots, k\}$, we aim to model the distribution $p(y | x; \theta)$

Given a training set { $(x^{(1)}, y^{(1)}), \dots$, we aim to model the distribution p(

$$\{(x^{(n)}, y^{(n)})\}, y^{(i)} \in \{1, 2, \cdots, k\},\ (y \mid x; \theta)$$

Categorical distribution, $p(y = k | x; \theta) = \phi_k$

s.t.
$$\sum_{i=1}^{k} \phi_i = 1$$

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 $\phi_i = \theta_i^T x$?



Softmax Function

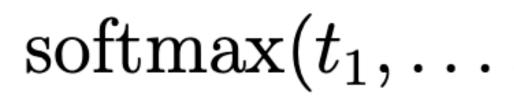


Softmax Function

Softmax: $\mathbb{R}^k \to \mathbb{R}^k$



Softmax:



Softmax Function

$$: \mathbb{R}^k \to \mathbb{R}^k$$

$$, t_k) = \begin{bmatrix} \frac{\exp(t_1)}{\sum_{j=1}^k \exp(t_j)} \\ \vdots \\ \frac{\exp(t_k)}{\sum_{j=1}^k \exp(t_j)} \end{bmatrix}$$



 $\operatorname{softmax}(t_1,\ldots)$

Softmax Function

Softmax: $\mathbb{R}^k \to \mathbb{R}^k$

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The denominator is a normalization constant

Let $(t_1, \ldots, t_k) = (\theta_1^\top x, \cdots, \theta_k^\top x)$

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$$\begin{bmatrix} P(y=1 \mid x; \theta) \\ \vdots \\ P(y=k \mid x; \theta) \end{bmatrix} = \operatorname{softmax}(t_1, \cdots, t_k) = \begin{bmatrix} \frac{\exp(\theta_1^\top x)}{\sum_{j=1}^k \exp(\theta_j^\top x)} \\ \vdots \\ \frac{\exp(\theta_k^\top x)}{\sum_{j=1}^k \exp(\theta_j^\top x)} \end{bmatrix}$$

 $, heta_k^ op x)$

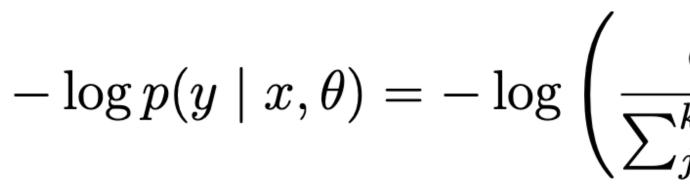
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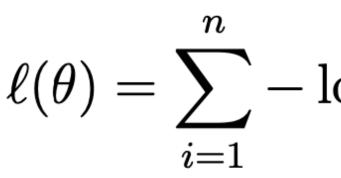
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$$P(y = i \mid x; \theta) = \phi_i = \frac{\exp(t_i)}{\sum_{j=1}^k \exp(t_j)} = \frac{\exp(\theta_i^\top x)}{\sum_{j=1}^k \exp(t_j)}$$

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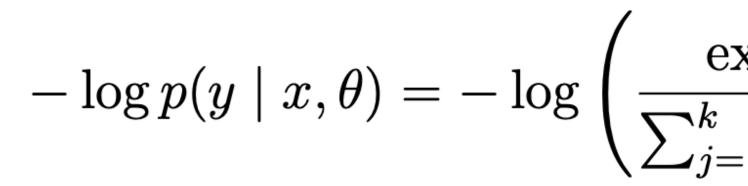
 $-\log p(y \mid x, \theta) = -\log \left(\frac{\exp(t_y)}{\sum_{i=1}^k \exp(t_i)} \right) = -\log \left(\frac{\exp(\theta_y^\top x)}{\sum_{i=1}^k \exp(\theta_i^\top x)} \right)$

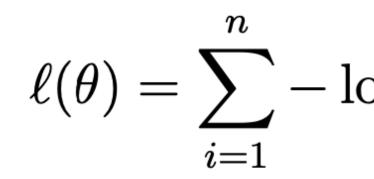




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 $\ell(\theta) = \sum_{i=1}^{n} -\log\left(\frac{\exp(\theta_{y^{(i)}}^{\top} x^{(i)})}{\sum_{i=1}^{k} \exp(\theta_{i}^{\top} x^{(i)})}\right)$



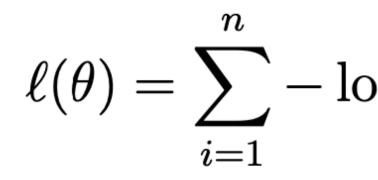


$$\frac{\exp(t_y)}{\sum_{j=1}^{k} \exp(t_j)} = -\log\left(\frac{\exp(\theta_y^{\top} x)}{\sum_{j=1}^{k} \exp(\theta_j^{\top} x)}\right)$$

$$\operatorname{og}\left(\frac{\exp(\theta_{y^{(i)}}^{\top}x^{(i)})}{\sum_{j=1}^{k}\exp(\theta_{j}^{\top}x^{(i)})}\right) \text{ Negative log likelihood}$$

Multi-Label Classification

$$-\log p(y \mid x, \theta) = -\log \left(\frac{\exp(t_y)}{\sum_{j=1}^k \exp(t_j)} \right) = -\log \left(\frac{\exp(\theta_y^\top x)}{\sum_{j=1}^k \exp(\theta_j^\top x)} \right)$$



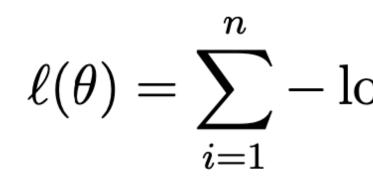
Cross-entropy loss $\ell_{ce} : \mathbb{R}^k \times \{$

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$$\{1,\ldots,k\} \to \mathbb{R}_{\geq 0}$$

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Cross-entropy loss
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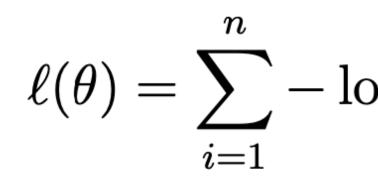
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Cross-entropy loss
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 $\ell_{ce}((t_1, \dots, t_k), y) = -\log\left(\frac{\exp(t_y)}{\sum_{j=1}^k \exp(t_j)}\right) \qquad \ell(\theta) = \sum_{i=1}^n \ell_{ce}((\theta_1^\top x^{(i)}, \dots, \theta_k^\top x^{(i)}), y^{(i)})$

$$\operatorname{og}\left(\frac{\exp(\theta_{y^{(i)}}^{\top}x^{(i)})}{\sum_{j=1}^{k}\exp(\theta_{j}^{\top}x^{(i)})}\right) \text{ Negative log likelihood}$$



 $\frac{\partial \ell_{\rm ce}(t, y)}{\partial t_i} = \phi_i - 1\{y = i\}$

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 $\frac{\partial \ell_{\rm ce}(t, y)}{\partial t_i} = \phi_i - 1\{y = i\}$

$$\phi_i = \frac{\exp(t_i)}{\sum_{j=1}^k \exp(t_j)}$$

$$\frac{\partial \ell_{\rm ce}(t, y)}{\partial t_i} = \phi_i - 1\{y = i\}$$

$\frac{\partial \ell_{ce}((\theta_1^{\top}x,\ldots,\theta_k^{\top}x),y)}{\partial \theta_i} = \frac{\partial \ell_{ce}((\theta_1^{\top}x,\ldots,\theta_k^{\top}x),y)}{\partial \theta_i}$

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$$\frac{\partial \ell(t, y)}{\partial t_i} \cdot \frac{\partial t_i}{\partial \theta_i} = (\phi_i - 1\{y = i\}) \cdot x$$

$$\frac{\partial \ell_{\rm ce}(t, y)}{\partial t_i} = \phi_i - 1\{y = i\}$$

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$$\frac{\partial \ell(\theta)}{\partial \theta_i} = \sum_{j=1}^n (\phi_i^{(j)} - 1\{y^{(j)} =$$

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 $i\})\cdot x^{(j)}$

$$\frac{\partial \ell_{\rm ce}(t, y)}{\partial t_i} = \phi_i - 1\{y = i\}$$

Chain rule $\frac{\partial \ell_{\rm ce}((\theta_1^\top x,\ldots,\theta_k^\top x),y)}{\partial \theta_i} = \frac{\partial \eta_i}{\partial \theta_i}$

$$\frac{\partial \ell(\theta)}{\partial \theta_i} = \sum_{j=1}^n (\phi_i^{(j)} - 1\{y^{(j)} =$$

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i}) · $x^{(j)}$ Intuitive explanation of the rule?



Given $f : \mathbb{R}^d \to \mathbb{R}$ find x s.t. f(x) = 0.

Given $f : \mathbb{R}^d \to \mathbb{R}$ find x s.t. f(x) = 0. $\nabla_{\theta} l(\theta) = 0$

Given $f : \mathbb{R}^d \to \mathbb{R}$ find x s.t. f(x) = 0. $\nabla_{\theta} l(\theta) = 0$

This is the update rule in 1d

 $x^{(t+1)} = x^{(t)} - \frac{f(x^{(t)})}{f'(x^{(t)})}$

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This is the update rule in 1d

 $x^{(t+1)} = x^{(t)} -$

$$\frac{f(x^{(t)})}{f'(x^{(t)})}$$

Solution to a linear equation $f'(x^{(t)})x^{(t+1)} + f(x^{(t)}) - x^{(t)}f'(x^{(t)}) = 0$

Given $f : \mathbb{R}^d \to \mathbb{R}$ find x s.t. f(x) = 0. $\nabla_{\theta} l(\theta) = 0$

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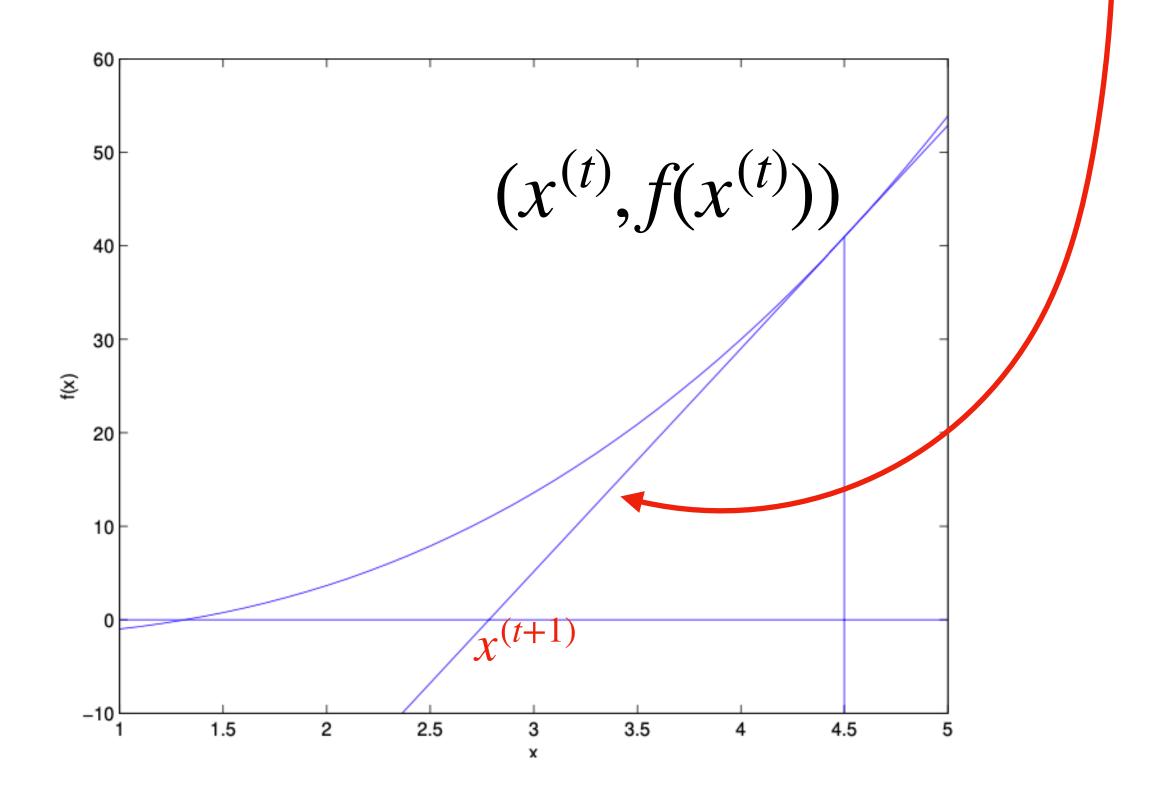
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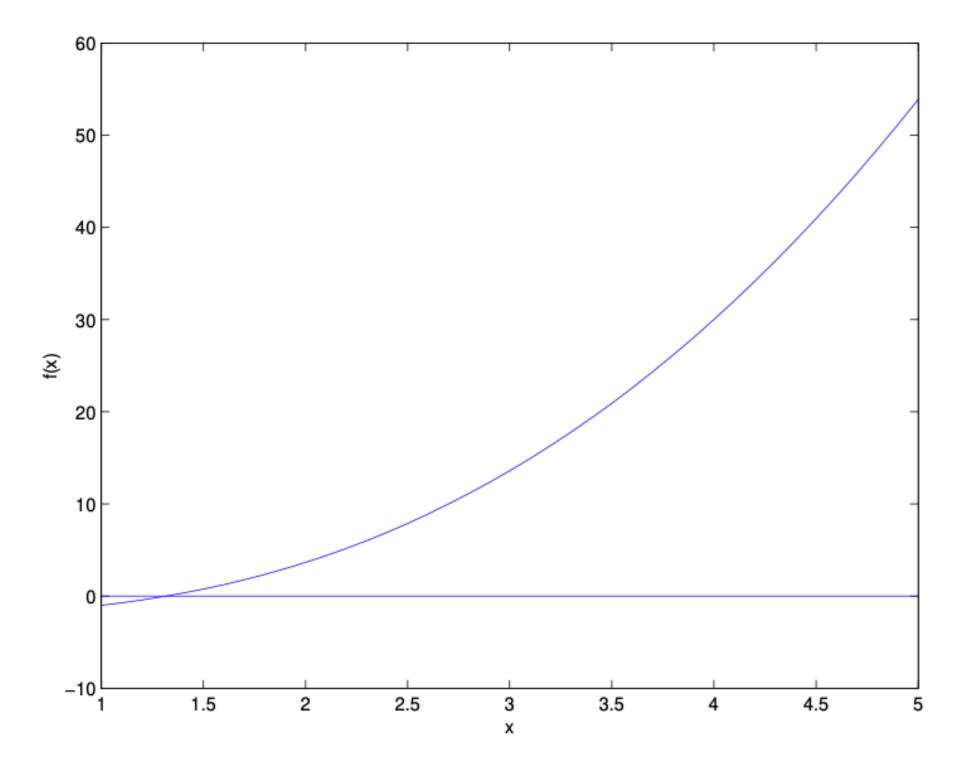
Solution to a linear equation $f'(x^{(t)})x^{(t+1)} + f(x^{(t)}) - x^{(t)}f'(x^{(t)}) = 0$

View it as a equation of $x^{(t+1)}$, and $x^{(t)}$ is a constant

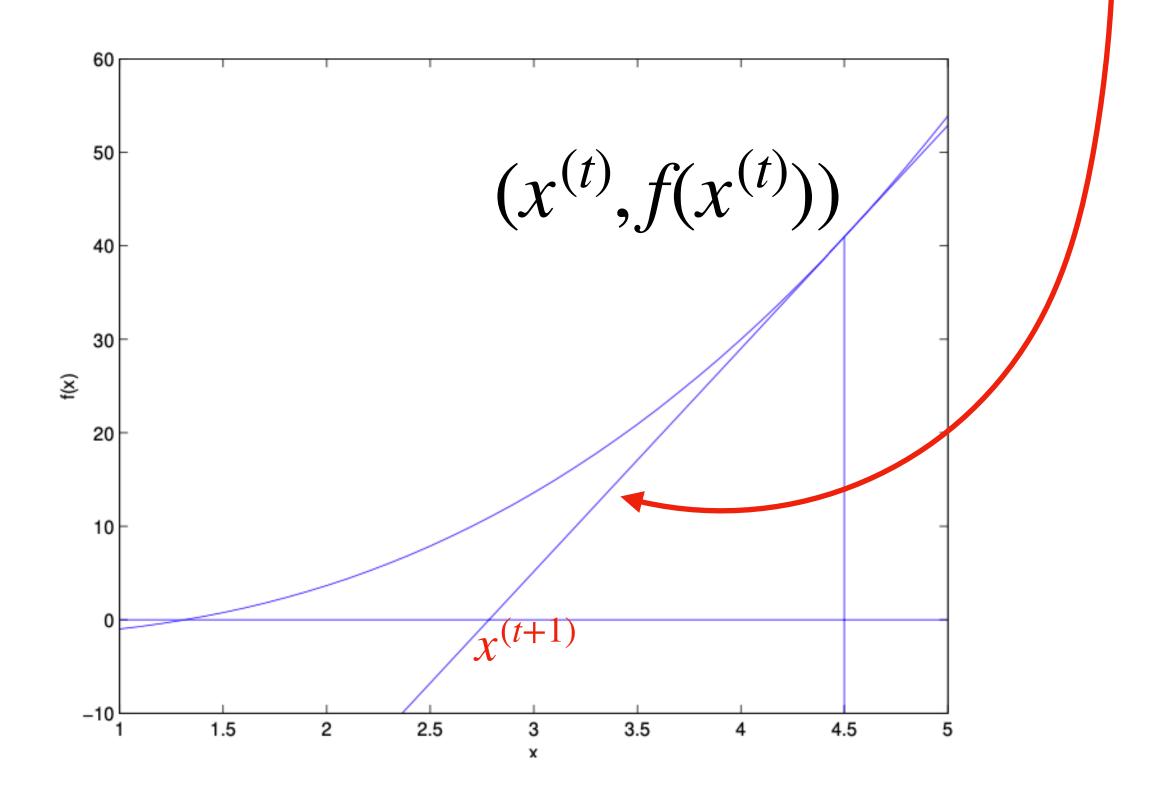
$$\frac{f(x^{(t)})}{f'(x^{(t)})}$$

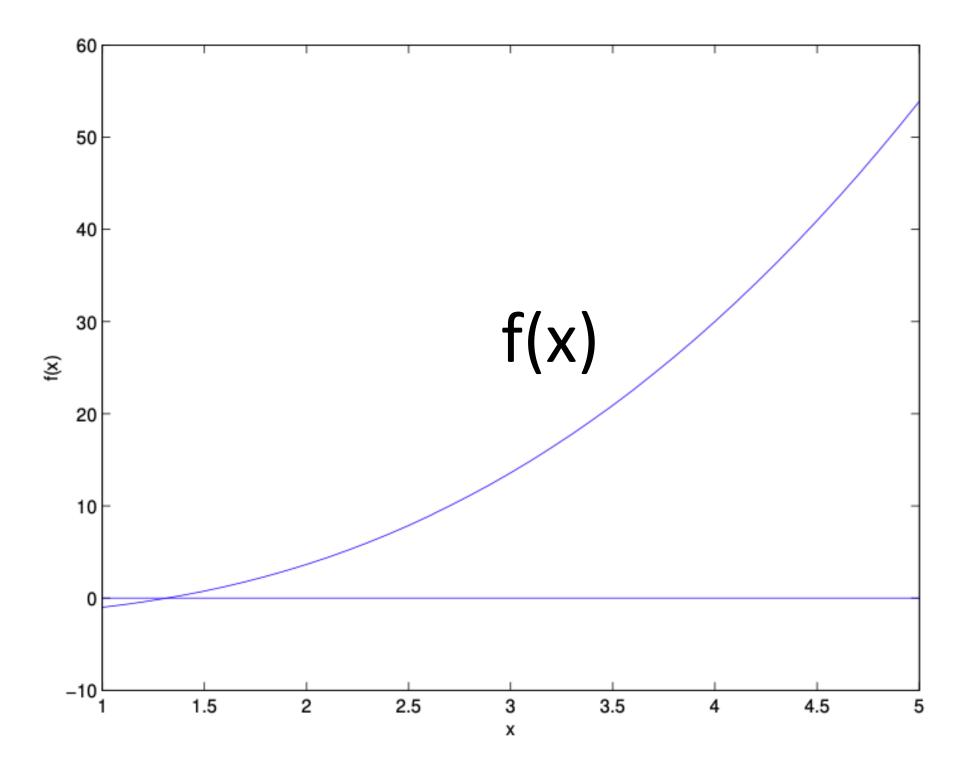
 $f'(x^{(t)})x + f(x^{(t)}) - x^{(t)}f'(x^{(t)}) = y$



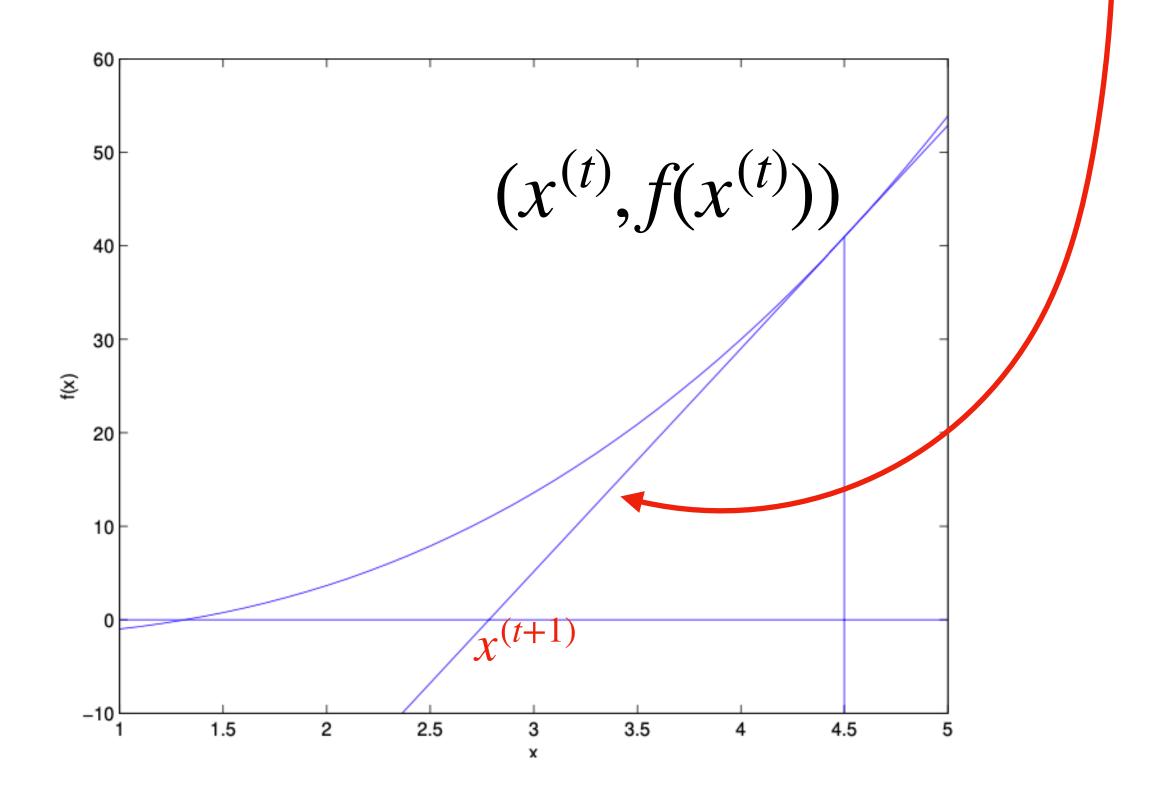


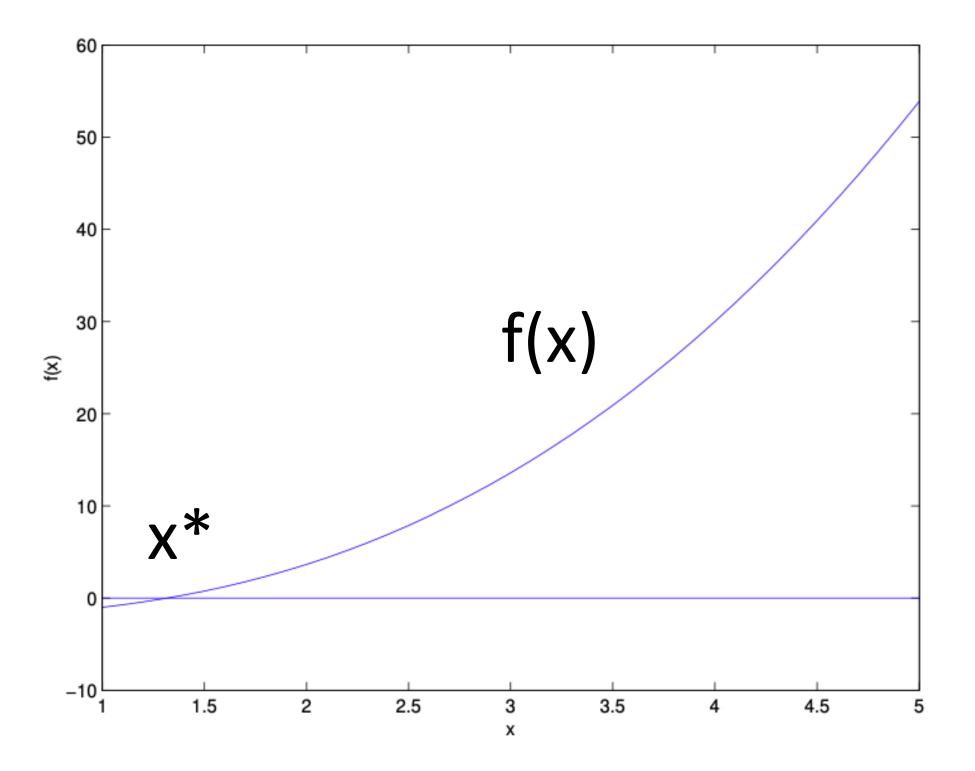
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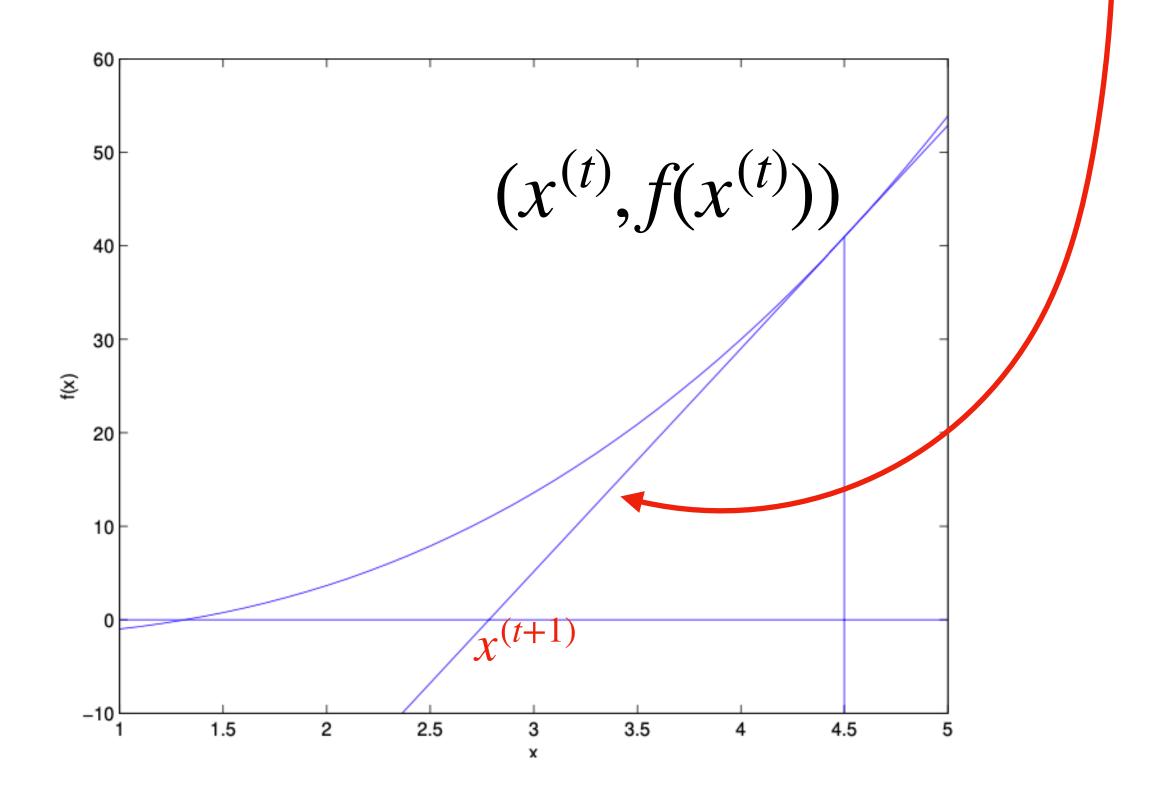


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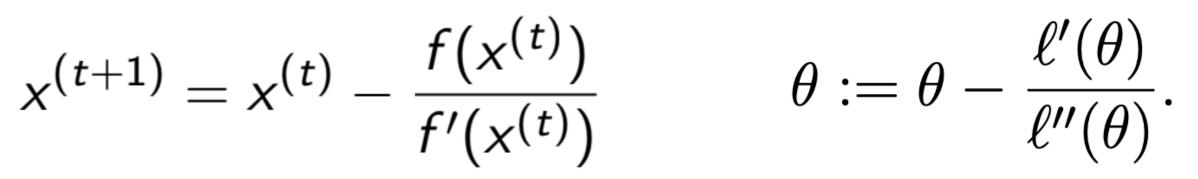


 $f'(x^{(t)})x + f(x^{(t)}) - x^{(t)}f'(x^{(t)}) = y$



Given $f : \mathbb{R}^d \to \mathbb{R}$ find x s.t. f(x) = 0. $\nabla_{\theta} l(\theta) = 0$

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Given $f : \mathbb{R}^d \to \mathbb{R}$ find x s.t. f(x) = 0. $\nabla_{\theta} l(\theta) = 0$

This is the update rule in 1d

 $x^{(t+1)} = x^{(t)} -$

$$-\frac{f(x^{(t)})}{f'(x^{(t)})} \qquad \qquad \theta := \theta - \frac{\ell'(\theta)}{\ell''(\theta)}.$$

It may converge very fast (quadratic local convergence!) Requires fewer iterations



Given $f : \mathbb{R}^d \to \mathbb{R}$ find x s.t. f(x) = 0. $\nabla_{\theta} l(\theta) = 0$

This is the update rule in 1d

 $x^{(t+1)} = x^{(t)} -$

For the likelihood, i.e., $f(\theta) = \nabla_{\theta} \ell(\theta)$ we need to generalize to a vector-valued function which has:

$$\theta^{(t+1)} = \theta^{(t)} - \left(H(\theta^{(t)})\right)^{-1} \nabla_{\theta} \ell(\theta^{(t)}).$$

in which
$$H_{i,j}(\theta) = \frac{\partial}{\partial \theta_i \partial \theta_j} \ell(\theta)$$
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When Newton's method is applied to maximize the logistic regression log likelihood function $I(\theta)$, the resulting method is also called Fisher scoring.













important models

Exponential Family

Exponential family unifies inference and learning for many



Rough Idea *"If P has a a special form, then inference and* learning come for free"

 $P(y;\eta) = b(y) e_{X}$

$$\operatorname{xp}\left\{\eta^{T}T(y)-a(\eta)
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Here y, $a(\eta)$, and b(y) are scalars. T(y) same dimension as η .

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 $P(y;\eta) = b(y) ex$

T(y) is called the **sufficient statistic**. b(y) is called the **base measure** – does *not* depend on η . $a(\eta)$ is called the **log partition function** – does *not* depend on y.

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$$1 = \sum_{y} P(y; \eta) = e^{-a(\eta)} \sum_{y} b(y) \exp\left\{\eta^{T} T(y)\right\}$$
$$\implies a(\eta) = \log \sum_{y} b(y) \exp\left\{\eta^{T} T(y)\right\}$$



(y)



Bernoulli random variable is an event (say flipping a coin) then:

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We need to show $a(\eta)$ is a function of $\log \frac{\phi}{1-\phi}$

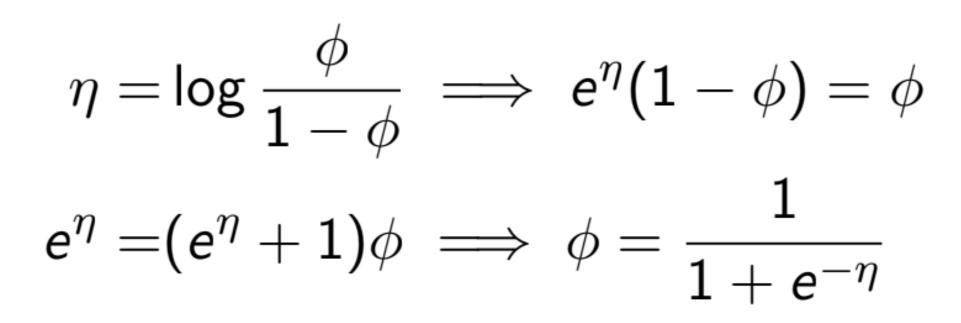
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We have verified Bernoulli distribution is in the exponential family

 $P(y; \mu) =$

Can we put it in the exponential family form?

 $P(y;\eta) =$

$$=\frac{1}{\sqrt{2\pi}}\exp\left\{-\frac{1}{2}(y-\mu)^2\right\}.$$

$$= b(y) \exp\left\{\eta^T T(y) - a(\eta)\right\}.$$

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 $\eta = \mu$

Notice that for a Gaussian with mean μ we had

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Is this true for general?

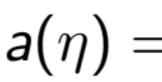
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Log Partition Function

Yes! Recall that



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Then, taking derivatives

 $\nabla_{\eta} a(\eta) = \frac{\sum_{y} T(y) b(y) \exp\left\{\eta^{T} T(y)\right\}}{\sum_{y} b(y) \exp\left\{\eta^{T} T(y)\right\}} = \mathbb{E}[T(y);\eta]$

 $a(\eta) = \log \sum_{y} b(y) \exp \left\{ \eta^T T(y) \right\}$

Many Other Exponential Models

- \blacktriangleright Binary \mapsto Bernoulli
- \blacktriangleright Multiple Classses \mapsto Multinomial
- \blacktriangleright Real \mapsto Gaussian
- \blacktriangleright Counts \mapsto Poisson
- \triangleright $\mathbb{R}_+ \mapsto$ Gamma, Exponential
- \blacktriangleright Distributions \mapsto Dirichlet

There are many canonical exponential family models:

Thank You! Q&A